## **Introduction to Machine Learning**

# **Gaussian Processes Bayesian Linear Model**

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#### MAP after observing 5 data points



#### **Learning goals**

- Know the Bayesian linear model
- The Bayesian LM returns a (posterior) distribution instead of a point estimate
- Know how to derive the posterior distribution for a Bayesian LM

Let  $\mathcal{D} = \left\{ (\bold{x}^{(1)}, y^{(1)}), ..., (\bold{x}^{(n)}, y^{(n)}) \right\}$  be a training set of i.i.d. observations from some unknown distribution.



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Let  $\mathbf{y} = (y^{(1)},...,y^{(n)})^\top$  and  $\mathbf{X} \in \mathbb{R}^{n \times p}$  be the design matrix where the i-th row contains vector **x** (*i*) .

The linear regression model is defined as

$$
y = f(\mathbf{x}) + \epsilon = \boldsymbol{\theta}^T \mathbf{x} + \epsilon
$$

or on the data:

$$
y^{(i)} = f\left(\mathbf{x}^{(i)}\right) + \epsilon^{(i)} = \boldsymbol{\theta}^T \mathbf{x}^{(i)} + \epsilon^{(i)}, \text{ for } i \in \{1, \ldots, n\}
$$

We now assume (from a Bayesian perspective) that also our parameter vector  $\theta$  is stochastic and follows a distribution. The observed values  $y^{(i)}$  differ from the function values  $f\left(\mathbf{x}^{(i)}\right)$  by some additive noise, which is assumed to be i.i.d. Gaussian

$$
e^{(i)} \sim \mathcal{N}(0, \sigma^2)
$$

 $\epsilon$ 

and independent of **x** and θ.

Let us assume we have **prior beliefs** about the parameter  $\theta$  that are represented in a prior distribution  $\boldsymbol{\theta} \sim \mathcal{N}(\mathbf{0}, \tau^2 \boldsymbol{I}_{\boldsymbol{\rho}}).$ 

Whenever data points are observed, we update the parameters' prior distribution according to Bayes' rule

$$
\underbrace{p(\boldsymbol{\theta}|\boldsymbol{\mathsf{X}}, \boldsymbol{\mathsf{y}})}_{\text{posterior}} = \frac{\overbrace{\boldsymbol{\rho}(\boldsymbol{\mathsf{y}}|\boldsymbol{\mathsf{X}}, \boldsymbol{\theta})}^{\text{likelihood}}\,\overbrace{q(\boldsymbol{\theta})}^{\text{prior}}}{\underbrace{\boldsymbol{\rho}(\boldsymbol{\mathsf{y}}|\boldsymbol{\mathsf{X}})}_{\text{marginal}}}.
$$

The posterior distribution of the parameter  $\theta$  is again normal distributed (the Gaussian family is self-conjugate):

$$
\boldsymbol{\theta} \mid \mathbf{X}, \mathbf{y} \sim \mathcal{N}(\sigma^{-2} \mathbf{A}^{-1} \mathbf{X}^{\top} \mathbf{y}, \mathbf{A}^{-1})
$$

with  $\mathbf{A} := \sigma^{-2} \mathbf{X}^\top \mathbf{X} + \frac{1}{\tau^2} \mathbf{I}_p$ .

**Note:** If the posterior distribution  $p(\theta | \mathbf{X}, \mathbf{y})$  are in the same probability distribution family as the prior  $q(\theta)$  w.r.t. a specific likelihood function  $p(\mathbf{y} \mid \mathbf{X}, \theta)$ , they are called **conjugate distributions**. The prior is then called a **conjugate prior** for the likelihood. The Gaussian family is self-conjugate: Choosing a Gaussian prior for a Gaussian Likelihood ensures that the posterior is Gaussian.



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MAP after observing 5 data points





MAP after observing 10 data points





MAP after observing 20 data points



#### **Proof:**

We want to show that

- for a Gaussian prior on  $\boldsymbol{\theta} \sim \mathcal{N}(\mathbf{0}, \tau^2 \boldsymbol{I}_{\rho})$
- for a Gaussian Likelihood  $\textsf{y} \mid \textsf{X}, \boldsymbol{\theta} \sim \mathcal{N}(\textsf{X}^\top \boldsymbol{\theta}, \sigma^2 \textsf{I}_n)$

the resulting posterior is Gaussian  $\mathcal{N}(\sigma^{-2} \mathbf{A}^{-1} \mathbf{X}^{\top} \mathbf{y}, \mathbf{A}^{-1})$  with  $\mathbf{A} := \sigma^{-2} \mathbf{X}^{\top} \mathbf{X} + \frac{1}{\tau^2} \mathbf{I}_p$ . Plugging in Bayes' rule and multiplying out yields

$$
\rho(\theta|\mathbf{X}, \mathbf{y}) \propto \rho(\mathbf{y}|\mathbf{X}, \theta) q(\theta) \propto \exp\left[-\frac{1}{2\sigma^2}(\mathbf{y} - \mathbf{X}\theta)^{\top}(\mathbf{y} - \mathbf{X}\theta) - \frac{1}{2\tau^2}\theta^{\top}\theta\right]
$$
\n
$$
= \exp\left[-\frac{1}{2}\left(\frac{\sigma^{-2}\mathbf{y}^{\top}\mathbf{y}}{\text{dosh}^2}\right) - 2\sigma^{-2}\mathbf{y}^{\top}\mathbf{X}\theta + \sigma^{-2}\theta^{\top}\mathbf{X}^{\top}\mathbf{X}\theta + \tau^{-2}\theta^{\top}\theta\right)\right]
$$
\n
$$
\propto \exp\left[-\frac{1}{2}\left(\sigma^{-2}\theta^{\top}\mathbf{X}^{\top}\mathbf{X}\theta + \tau^{-2}\theta^{\top}\theta - 2\sigma^{-2}\mathbf{y}^{\top}\mathbf{X}\theta\right)\right]
$$
\n
$$
= \exp\left[-\frac{1}{2}\theta^{\top}\left(\sigma^{-2}\mathbf{X}^{\top}\mathbf{X} + \tau^{-2}\mathbf{I}_{\rho}\right)\theta + \sigma^{-2}\mathbf{y}^{\top}\mathbf{X}\theta\right]
$$

This expression resembles a normal density - except for the term in red!

**Note:** We need not worry about the normalizing constant since its mere role is to convert probability functions to density functions with a total probability of one. We subtract a (not yet defined) constant *c* while compensating for this change by adding the respective terms ("adding 0"), emphasized in green:

$$
\rho(\theta|\mathbf{X}, \mathbf{y}) \propto \exp\left[-\frac{1}{2}(\theta - c)^{\top} \mathbf{A}(\theta - c) - c^{\top} \mathbf{A}\theta + \underbrace{\frac{1}{2}c^{\top} \mathbf{A}c}_{\text{doesn't depend on }\theta} + \sigma^{-2} \mathbf{y}^{\top} \mathbf{X}\theta\right]
$$
\n
$$
\propto \exp\left[-\frac{1}{2}(\theta - c)^{\top} \mathbf{A}(\theta - c) - c^{\top} \mathbf{A}\theta + \sigma^{-2} \mathbf{y}^{\top} \mathbf{X}\theta\right]
$$

If we choose  $c$  such that  $-c^\top {\bf A}\theta + \sigma^{-2} {\bf y}^\top {\bf X}\theta = 0,$  the posterior is normal with mean  $c$ and covariance matrix **A**<sup>−1</sup>. Taking into account that **A** is symmetric, this is if we choose

$$
\sigma^{-2} \mathbf{y}^{\top} \mathbf{X} = \mathbf{c}^{\top} \mathbf{A}
$$

$$
\Leftrightarrow \quad \sigma^{-2} \mathbf{y}^{\top} \mathbf{X} \mathbf{A}^{-1} = \mathbf{c}^{\top}
$$

$$
\Leftrightarrow \quad \mathbf{c} = \sigma^{-2} \mathbf{A}^{-1} \mathbf{X}^{\top} \mathbf{y}
$$

as claimed.

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Based on the posterior distribution

$$
\boldsymbol{\theta} \mid \mathbf{X}, \mathbf{y} \sim \mathcal{N}(\sigma^{-2} \mathbf{A}^{-1} \mathbf{X}^{\top} \mathbf{y}, \mathbf{A}^{-1})
$$

we can derive the predictive distribution for a new observations **x**∗. The predictive distribution for the Bayesian linear model, i.e. the distribution of  $\boldsymbol{\theta}^{\top}\mathbf{x}_{*}$ , is

$$
y_*\mid \textbf{X},\textbf{y},\textbf{x}_* \sim \mathcal{N}(\sigma^{-2}\textbf{y}^\top\textbf{X}\textbf{A}^{-1}\textbf{x}_*,\textbf{x}_*^\top\textbf{A}^{-1}\textbf{x}_*)
$$

(applying the rules for linear transformations of Gaussians).



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For every test input **x**∗, we get a distribution over the prediction *y*∗. In particular, we get a posterior mean (orange) and a posterior variance (grey region equals  $+/-$  two times standard deviation).



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## **SUMMARY: THE BAYESIAN LINEAR MODEL**

- By switching to a Bayesian perspective, we do not only have point estimates for the parameter  $\theta$ , but whole **distributions**
- **•** From the posterior distribution of  $\theta$ , we can derive a predictive distribution for  $y_* = \theta^\top \mathbf{x}_*.$
- We can perform online updates: Whenever datapoints are observed, we can update the **posterior distribution** of θ

Next, we want to develop a theory for general shape functions, and not only for linear function.