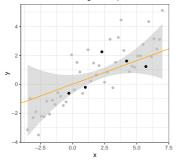
Introduction to Machine Learning

Gaussian Processes Bayesian Linear Model

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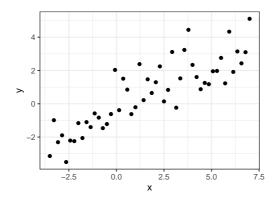
MAP after observing 5 data points



Learning goals

- Know the Bayesian linear model
- The Bayesian LM returns a (posterior) distribution instead of a point estimate
- Know how to derive the posterior distribution for a Bayesian LM

Let $\mathcal{D} = \{(\mathbf{x}^{(1)}, y^{(1)}), ..., (\mathbf{x}^{(n)}, y^{(n)})\}$ be a training set of i.i.d. observations from some unknown distribution.



Let $\mathbf{y} = (y^{(1)}, ..., y^{(n)})^{\top}$ and $\mathbf{X} \in \mathbb{R}^{n \times p}$ be the design matrix where the i-th row contains vector $\mathbf{x}^{(i)}$.

The linear regression model is defined as

$$y = f(\mathbf{x}) + \epsilon = \boldsymbol{\theta}^T \mathbf{x} + \epsilon$$

or on the data:

$$y^{(i)} = f\left(\mathbf{x}^{(i)}\right) + \epsilon^{(i)} = \boldsymbol{\theta}^T \mathbf{x}^{(i)} + \epsilon^{(i)}, \text{ for } i \in \{1, \dots, n\}$$

We now assume (from a Bayesian perspective) that also our parameter vector $\boldsymbol{\theta}$ is stochastic and follows a distribution. The observed values $y^{(i)}$ differ from the function values $f(\mathbf{x}^{(i)})$ by some additive noise, which is assumed to be i.i.d. Gaussian

$$\epsilon^{(i)} \sim \mathcal{N}(\mathbf{0}, \sigma^2)$$

f

and independent of \mathbf{x} and $\boldsymbol{\theta}$.

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Let us assume we have **prior beliefs** about the parameter θ that are represented in a prior distribution $\theta \sim \mathcal{N}(\mathbf{0}, \tau^2 \mathbf{I}_p)$.

Whenever data points are observed, we update the parameters' prior distribution according to Bayes' rule

$$\underbrace{p(\boldsymbol{\theta}|\mathbf{X},\mathbf{y})}_{\text{posterior}} = \frac{\overbrace{p(\mathbf{y}|\mathbf{X},\boldsymbol{\theta})}^{\text{likelihood}} \overbrace{q(\boldsymbol{\theta})}^{\text{prior}}}{\underbrace{p(\mathbf{y}|\mathbf{X})}_{\text{marginal}}}.$$

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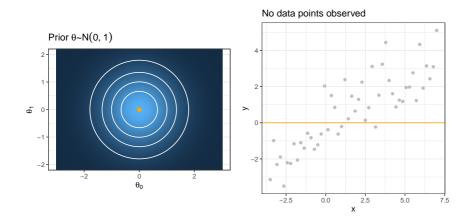
The posterior distribution of the parameter θ is again normal distributed (the Gaussian family is self-conjugate):

$$oldsymbol{ heta} \mid \mathbf{X}, \mathbf{y} \sim \mathcal{N}(\sigma^{-2} \mathbf{A}^{-1} \mathbf{X}^{ op} \mathbf{y}, \mathbf{A}^{-1})$$

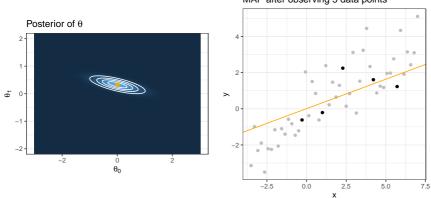
with $\mathbf{A} := \sigma^{-2} \mathbf{X}^{\top} \mathbf{X} + \frac{1}{\tau^2} \mathbf{I}_p$.

Note: If the posterior distribution $p(\theta \mid \mathbf{X}, \mathbf{y})$ are in the same probability distribution family as the prior $q(\theta)$ w.r.t. a specific likelihood function $p(\mathbf{y} \mid \mathbf{X}, \theta)$, they are called **conjugate distributions**. The prior is then called a **conjugate prior** for the likelihood. The Gaussian family is self-conjugate: Choosing a Gaussian prior for a Gaussian Likelihood ensures that the posterior is Gaussian.

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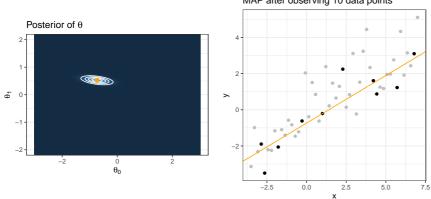


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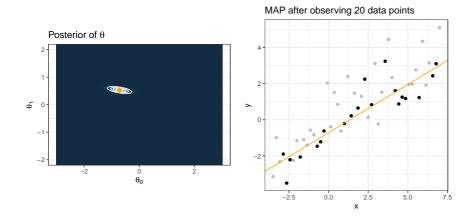
MAP after observing 5 data points





MAP after observing 10 data points





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Proof:

We want to show that

- for a Gaussian prior on $\boldsymbol{\theta} \sim \mathcal{N}(\mathbf{0}, \tau^2 \boldsymbol{I}_p)$
- for a Gaussian Likelihood $y \mid \mathbf{X}, \boldsymbol{\theta} \sim \mathcal{N}(\mathbf{X}^{\top}\boldsymbol{\theta}, \sigma^2 \boldsymbol{I}_n)$

the resulting posterior is Gaussian $\mathcal{N}(\sigma^{-2}\mathbf{A}^{-1}\mathbf{X}^{\top}\mathbf{y}, \mathbf{A}^{-1})$ with $\mathbf{A} := \sigma^{-2}\mathbf{X}^{\top}\mathbf{X} + \frac{1}{\tau^{2}}\mathbf{I}_{\rho}$. Plugging in Bayes' rule and multiplying out yields

$$p(\theta | \mathbf{X}, \mathbf{y}) \propto p(\mathbf{y} | \mathbf{X}, \theta) q(\theta) \propto \exp\left[-\frac{1}{2\sigma^2} (\mathbf{y} - \mathbf{X}\theta)^\top (\mathbf{y} - \mathbf{X}\theta) - \frac{1}{2\tau^2} \theta^\top \theta\right]$$

$$= \exp\left[-\frac{1}{2} \left(\underbrace{\sigma^{-2} \mathbf{y}^\top \mathbf{y}}_{\text{doesn't depend on } \theta} - 2\sigma^{-2} \mathbf{y}^\top \mathbf{X}\theta + \sigma^{-2} \theta^\top \mathbf{X}^\top \mathbf{X}\theta + \tau^{-2} \theta^\top \theta\right)\right]$$

$$\propto \exp\left[-\frac{1}{2} \left(\sigma^{-2} \theta^\top \mathbf{X}^\top \mathbf{X}\theta + \tau^{-2} \theta^\top \theta - 2\sigma^{-2} \mathbf{y}^\top \mathbf{X}\theta\right)\right]$$

$$= \exp\left[-\frac{1}{2} \theta^\top \underbrace{\left(\sigma^{-2} \mathbf{X}^\top \mathbf{X} + \tau^{-2} \mathbf{I}_{\rho}\right)}_{\mathbf{x} - \mathbf{x}} \theta + \sigma^{-2} \mathbf{y}^\top \mathbf{X}\theta\right]$$

This expression resembles a normal density - except for the term in red!

 \mathbf{x}

Note: We need not worry about the normalizing constant since its mere role is to convert probability functions to density functions with a total probability of one. We subtract a (not yet defined) constant *c* while compensating for this change by adding the respective terms ("adding 0"), emphasized in green:

$$p(\theta|\mathbf{X}, \mathbf{y}) \propto \exp\left[-\frac{1}{2}(\theta - c)^{\top} \mathbf{A}(\theta - c) - c^{\top} \mathbf{A}\theta + \underbrace{\frac{1}{2}c^{\top} \mathbf{A}c}_{\text{doesn't depend on } \theta} + \sigma^{-2} \mathbf{y}^{\top} \mathbf{X}\theta\right]$$
$$\propto \exp\left[-\frac{1}{2}(\theta - c)^{\top} \mathbf{A}(\theta - c) - c^{\top} \mathbf{A}\theta + \sigma^{-2} \mathbf{y}^{\top} \mathbf{X}\theta\right]$$

If we choose *c* such that $-c^{\top} \mathbf{A} \theta + \sigma^{-2} \mathbf{y}^{\top} \mathbf{X} \theta = 0$, the posterior is normal with mean *c* and covariance matrix \mathbf{A}^{-1} . Taking into account that **A** is symmetric, this is if we choose

$$\sigma^{-2} \mathbf{y}^{\mathsf{T}} \mathbf{X} = \boldsymbol{c}^{\mathsf{T}} \mathbf{A}$$
$$\Leftrightarrow \quad \sigma^{-2} \mathbf{y}^{\mathsf{T}} \mathbf{X} \mathbf{A}^{-1} = \boldsymbol{c}^{\mathsf{T}}$$
$$\Leftrightarrow \quad \boldsymbol{c} = \sigma^{-2} \mathbf{A}^{-1} \mathbf{X}^{\mathsf{T}} \mathbf{y}$$

as claimed.

Based on the posterior distribution

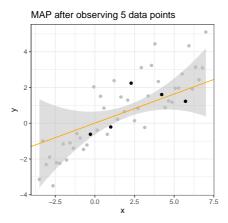
$$oldsymbol{ heta} \mid \mathbf{X}, \mathbf{y} \sim \mathcal{N}(\sigma^{-2} oldsymbol{A}^{-1} \mathbf{X}^{ op} \mathbf{y}, oldsymbol{A}^{-1})$$

we can derive the predictive distribution for a new observations \mathbf{x}_* . The predictive distribution for the Bayesian linear model, i.e. the distribution of $\boldsymbol{\theta}^{\top}\mathbf{x}_*$, is

$$y_* \mid \mathbf{X}, \mathbf{y}, \mathbf{x}_* \sim \mathcal{N}(\sigma^{-2} \mathbf{y}^\top \mathbf{X} \mathbf{A}^{-1} \mathbf{x}_*, \mathbf{x}_*^\top \mathbf{A}^{-1} \mathbf{x}_*)$$

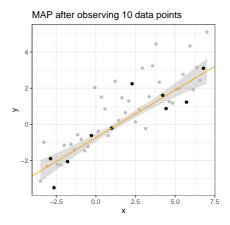
(applying the rules for linear transformations of Gaussians).

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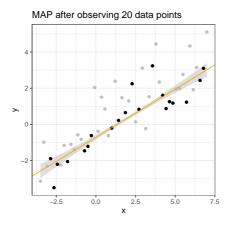


For every test input \mathbf{x}_* , we get a distribution over the prediction y_* . In particular, we get a posterior mean (orange) and a posterior variance (grey region equals +/- two times standard deviation).



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SUMMARY: THE BAYESIAN LINEAR MODEL

- By switching to a Bayesian perspective, we do not only have point estimates for the parameter θ , but whole **distributions**
- From the posterior distribution of θ, we can derive a predictive distribution for y_{*} = θ^Tx_{*}.
- We can perform online updates: Whenever datapoints are observed, we can update the **posterior distribution** of θ

Next, we want to develop a theory for general shape functions, and not only for linear function.

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