# **Algorithms and Data Structures**

**Quadrature Laplace's method**

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#### **Learning goals**

**•** Laplace's method

**Target:** Approximate integral of function *f* with the following properties:

- The mass concentrates on a small area around a center and the function has very rapidly decreasing tails ("similarity" to the density of a normal distribution)
- The function we want to integrate is the density of a random variable that is approximately normally distributed

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In particular, we assume that *f*

- Can only be positive
- Is two times continuously differentiable
- $\bullet$  Has a **global maximum** at  $x_0$

We could approximate the area underneath the graph of the function with a staircase function and represent the integral with a very simple formula that depends on  $f(x_0)$ :

$$
\int f(x) \ dx \approx f(x_0) \cdot c
$$



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But instead of the step function we would like to choose a function that approximates *f* **better** and which has well-known properties.

**Idea:** Approximate the integral using the density function of the normal distribution!

**How?** We center and scale the density function of the normal distribution such that it approximates *f* "best possible".

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In other words: We determine **expectation** and **standard deviation** of a normal distribution such that the corresponding density function fits best possible to the function *f* we are interested in.

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#### **Mathematical derivation:**

Let there be a function  $f$  with a global maximum at  $x_0$ .

We define  $h(x) := \log f(x)$  as the logarithmized function and rewrite the integral

$$
\int_a^b f(x) \ dx = \int_a^b \exp(\underbrace{\log f(x)}_{:=h(x)}) \ dx
$$

Using Taylor's theorem around  $x_0$  we obtain

$$
\int_{a}^{b} \exp(h(x)) \approx \int_{a}^{b} \exp\left(h(x_{0}) + h'(x_{0})(x - x_{0}) + \frac{1}{2}h''(x_{0})(x - x_{0})^{2}\right) dx
$$



 $x_0$  is also the maximum of  $h(x) = \log(f(x))$  . Hence,  $h'(x_0) = 0$  and the second summand disappears:

$$
\int_a^b \exp\left(h(x)\right) \ dx \approx \int_a^b \exp\left(h(x_0) + \frac{1}{2}h''(x_0)(x-x_0)^2\right) dx
$$

We take advantage of the fact that  $exp(x + y) = exp(x) exp(y)$ 

$$
\int_a^b \exp\left(h(x_0)\right) \cdot \exp\left(\frac{1}{2}h''(x_0)(x-x_0)^2\right) dx
$$

and pull the constant  $\exp(h(x_0))$  out of the integral

$$
\exp\left(h(x_0)\right)\cdot\int_a^b \exp\left(\frac{1}{2}h''(x_0)(x-x_0)^2\right)dx
$$



Within the integral there is now an expression which "almost" corresponds to the density of a normal distribution with expectation  $\mu:=x_0$  and variance  $\sigma^2:=-h''(x_0)^{-1}$ :

$$
\int_{a}^{b} f(x)dx \approx \exp(h(x_0)) \cdot \int_{a}^{b} \exp\left(\frac{1}{2}h''(x_0)(x-x_0)^2\right) dx
$$

$$
= \exp(h(x_0)) \cdot \int_{a}^{b} \exp\left(-\frac{1}{2}\frac{(x-x_0)^2}{-h''(x_0)^{-1}}\right) dx
$$

$$
= \exp(h(x_0)) \cdot \int_{a}^{b} \exp\left(-\frac{1}{2}\frac{(x-\mu)^2}{\sigma^2}\right) dx
$$

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 $-h''(x_0)^{-1}$  must be truly positive to correspond to the variance of a normal distribution. Since  $h(x)$  has a global maximum in  $x<sub>0</sub>$ , the second derivative at this point is negative and therefore  $-h''(x_0)^{-1}>0.$ 

If we add (and cancel) the multiplicative constant  $c = \frac{1}{\sqrt{2}}$  $\frac{1}{2\pi\sigma^2}$ , we obtain

$$
\int_{a}^{b} f(x)dx \approx \frac{1}{c} \cdot \exp(h(x_{0})) \cdot \int_{a}^{b} \underbrace{c \cdot \exp\left(-\frac{1}{2}\frac{(x-\mu)^{2}}{\sigma^{2}}\right)}_{\text{Density ND}} dx
$$
\n
$$
= \frac{1}{c} \underbrace{\exp(h(x_{0}))}_{f(x_{0})} \cdot \int_{a}^{b} \phi_{\mu,\sigma^{2}}(x) dx
$$
\n
$$
= \frac{1}{c} f(x_{0}) \cdot (\Phi_{\mu,\sigma^{2}}(b) - \Phi_{\mu,\sigma^{2}}(a))
$$

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where  $\phi_{\mu,\sigma^2}(x)$  denotes the density and  $\Phi_{\mu,\sigma^2}(x)$  the distribution function of a normal distribution with expectation  $\mu$  and variance  $\sigma^2.$ 

For integration limits  $b = \infty$  and  $a = -\infty$  Laplace's method of *f* is then

$$
\int_{-\infty}^{\infty} f(x) dx \approx \frac{1}{c} \cdot f(x_0) \cdot (\Phi_{\mu,\sigma^2}(+\infty) - \Phi_{\mu,\sigma^2}(-\infty))
$$

$$
= \sqrt{-\frac{2\pi}{h''(x_0)}} \cdot f(x_0)
$$

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$$

with  $h(x) = \log f(x)$ .

Laplace's method thus corresponds to a value that only depends on the maximum of the function  $f(x_0)$  and the curvature of the logarithmic function  $h''(x_0)$ .

Laplace's method also works well in higher dimensions. For  $f: \mathbb{R}^m \to \mathbb{R}$  with global maximum in  $x_0$  the generalized form is given by

$$
I(f) \approx (2\pi)^{m/2} \det(-H_f(x_0)^{-1})^{1/2} \exp(f(x_0))
$$

where  $H_f(x_0)$  denotes the Hessian matrix of f at  $x_0$ . Since  $x_0$  is a global maximum,  $H_f(x_0)$  is negative definite.

The problem of integration is reduced to

- Solving an optimization problem  $\rightarrow$  find  $x_0$
- Determining the second derivative  $h''(x)$  (or generally the Hessian matrix  $H_f(\mathbf{x})$  at the optimal position  $x_0$ .

Instead of integration, an optimization problem must now be solved, which is often much easier and faster.



**Application example:** Bayesian computation

**Given**:

 $x|\lambda \sim$  Poisson( $\lambda$ ) (Likelihood)  $\lambda \sim$  Gamma $(\alpha, \beta)$  (Prior)

**Wanted**: Posterior density of the parameter λ given *n* observations  $\boldsymbol{x} = (x^{(1)}, x^{(2)}, ..., x^{(n)})$ 

$$
\mathit{Posterior}\atop \mathit{p}(\lambda|\mathbf{x}) = \frac{\mathit{Likelihood}\atop \mathit{p}(\mathbf{x}|\lambda)\cdot\pi(\lambda)}{\int\mathit{p}(\mathbf{x}|\lambda)\cdot\pi(\lambda)\;d\lambda}
$$

The density of the gamma distribution is given by  $\pi_{\alpha,\beta}(\lambda)=\frac{1}{\beta^\alpha\Gamma(\alpha)}\lambda^{\alpha-1}\exp(-\lambda\beta)$ 





To keep the calculations simple, we calculate the posterior density for only **one** observation *x*.

The posterior density of  $\lambda$  given the observation x is (except for one constant)

$$
p(\lambda|x) \propto \lambda^{x+\alpha-1} \exp\left(-\frac{\lambda}{1/\beta+1}\right) =: f(\lambda).
$$

So to determine the posterior density  $p(\lambda|x)$  exactly, we search for the normalization constant  $c$ , which ensures that  $\int c \cdot f(\lambda) \ d\lambda =$  1, hence

$$
c \cdot \int f(\lambda) d\lambda = 1
$$
  

$$
c = \frac{1}{\int f(\lambda) d\lambda}
$$



**Goal:** Approximation of  $\int f(\lambda) d\lambda$  with  $f(\lambda) = \lambda^{x+\alpha-1} \exp \left(-\frac{\lambda}{1/\beta} \right)$  $\frac{\lambda}{1/\beta+1}$ 

We calculate  $h(\lambda) = \log f(\lambda)$ 

$$
h(\lambda) = \log \left( \lambda^{x+\alpha-1} \cdot \exp(-\frac{\lambda}{1/\beta+1}) \right)
$$
  
=  $(x+\alpha-1)\log \lambda - \frac{\lambda}{1/\beta+1}$   

$$
h'(\lambda) = \frac{x+\alpha-1}{\lambda} - \frac{1}{1/\beta+1}
$$
  

$$
h''(\lambda) = -\frac{x+\alpha-1}{\lambda^2}
$$

$$
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$$

To approximate the integral using Laplace's method, we need  $\lambda_0:=$  arg max  $f(\lambda)$  and  $h''(\lambda_0),$  where  $h(\lambda):=\log(f(\lambda)).$ 

The maximum of  $f(\lambda)$  is the same as the maximum of  $h(\lambda)$  (easier to calculate)

$$
\frac{x+\alpha-1}{\lambda} - \frac{1}{1/\beta+1} = 0
$$
  

$$
\lambda_0 = \frac{x+\alpha-1}{1/\beta+1}
$$

and thus

$$
h''(\lambda_0) = -\frac{(1/\beta + 1)^2}{x + \alpha - 1}
$$

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We insert  $\lambda_{0}=\frac{x+\alpha-1}{1/\beta+1}$  $\frac{\alpha+\alpha-1}{1/\beta+1}$  and  $\mathit{h}''(\lambda_0)$  into the formula for Laplace's method and obtain

$$
\int f(\lambda) d\lambda \approx \sqrt{-\frac{2\pi}{h''(\lambda_0)}} \cdot f(\lambda_0)
$$

$$
= \sqrt{2\pi} \cdot \frac{\sqrt{x + \alpha - 1}}{1/\beta + 1} \cdot f(\lambda_0)
$$

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Hence, the normalization constant *c* can be approximated by

$$
c = \frac{1}{\int f(\lambda) d\lambda} \approx \frac{1}{\sqrt{2\pi}} \frac{1/\beta + 1}{\sqrt{x + \alpha - 1}} \cdot \frac{1}{f(\lambda_0)}
$$

When calculating posterior distributions, Laplace's method provides a good approximation if

- The number *n* of observations is large
- The posterior distributions are roughly symmetric

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